

More Giving or More Givers?

The Effects of Tax Incentives on Charitable Donations in the UK

Miguel Almunia, Ben Lockwood and Kim Scharf

University of Warwick

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- How should govts support the private provision of public goods?
- Sensitivity of donations (g) to changes in the price of giving relative to consumption (p) is a key element to answer this question:

$$\varepsilon_{g,p} = \frac{\partial g(p, y)}{\partial p}$$

- Many estimates from different countries – mainly the US
 - Multiple empirical approaches: diff-in-diff, IV, quantile regs
 - Focus on intensive margin, ignoring extensive-margin decision
 - “Consensus” US estimate $\varepsilon_{g,p} \approx -1$ (eg, Bakija & Heim, 2011)
 - For France, Fack & Landais (2010) obtain $\varepsilon_{g,p} \in (-0.6, -0.2)$

- Policy change: major UK income tax reform in 2010
 - Two new tax brackets at the top: $\tau = 50\%$ (= 60% for short range)
 - Price of giving relative to consumption: $p_{it} = 1 - \tau_{it}$, as in the US
- Data: new administrative dataset of UK income tax returns
 - Universe of self-assessment taxpayers: $N = 75$ million
 - Period 2004/05 through 2012/13
- Estimation techniques:
 - Separate estimation of intensive and extensive-margin elasticities
 - New IV strategy to deal with endogenous earnings responses to reform
 - Pseudo-Poisson regressions to estimate total elasticity
 - Two-step model to account for sample selection

- Separate estimation of intensive and extensive-margin responses
 - Infeasible in US studies because the decision to itemize is endogenous (standard deduction system)
- First UK estimates of the price elasticity of giving using admin data
 - Existing estimates from donors' aggregate data (Khanna et al 1995) or lab experiments (Scharf and Smith 2014)
 - Relevant for public policy debate in the UK: charities' annual income approx. £60 billion
- New theoretical framework to derive policy implications of the price elasticity
 - Beyond the simple “unit-elasticity” rule

Institutional Context: Gift Aid

- 1 **Match:** charity collects donations from individual donors and HM Revenue and Customs (HMRC) matches those at the basic marginal income tax rate (τ_b):

$$P_b = 1 - \tau_b$$

- 2 **Rebate:** higher-rate taxpayers filing a 'self-assessment' return can also claim deduction on the difference between higher and basic marginal tax rates ($\tau_h - \tau_b$):

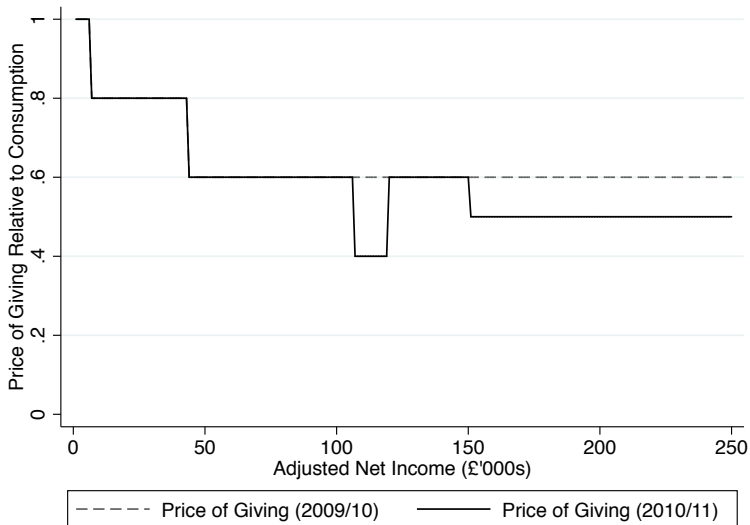
$$P_h = 1 - \tau_h$$

- In sum, Gift Aid incentives are economically equivalent to the US-style deduction system

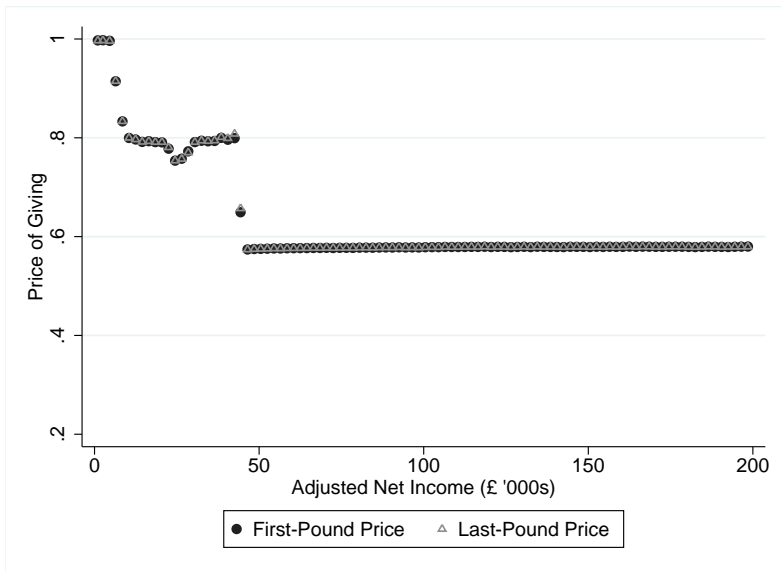
Gift Aid: Example

- A basic rate taxpayer makes a £100 donation to charity out of after-tax income:
 - Charity claims additional £25 from HMRC ($\tau_b = 20\%$)
 - Gross donation (D): £125 $\left(= 100 \times \frac{1}{1-\tau_b}\right)$
- A higher-rate taxpayer makes a £100 donation to charity out of after-tax income
 - Charity claims additional £25 from HMRC
 - Gross donation (D): £125
 - Taxpayer deducts £25 from income tax ($\tau_h = 40\%$)
 - Net donation: £75

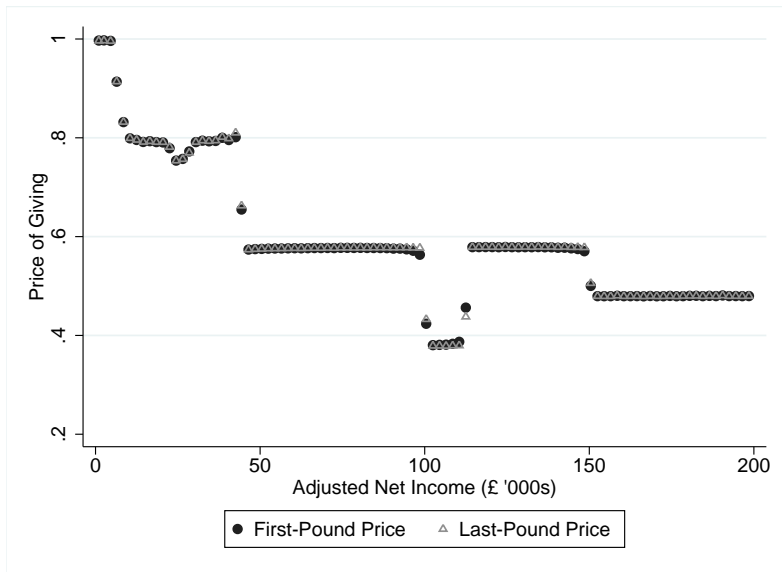
UK income tax reform of 2010



Tax-Price of Giving in the Data (2009/10)



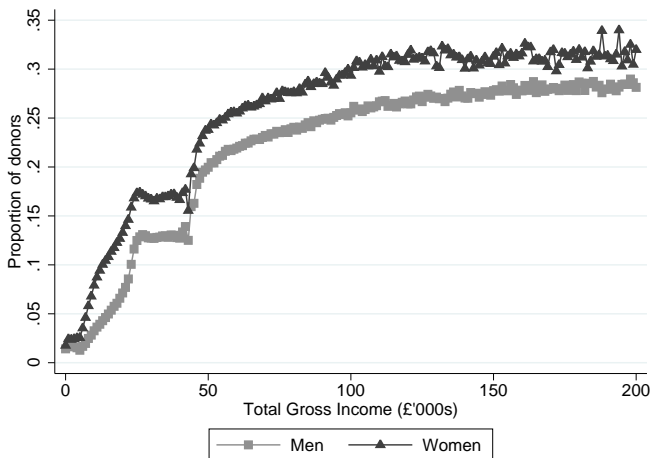
Tax-Price of Giving in the Data (2010/11)



- New panel of administrative data obtained from HMRC
- Full population of Self-Assessment (SA) Income Tax returns for the period 2004/05 through 2012/13
 - Approximately 8-9 million returns *per year*: $N = 75$ million
 - Does not include about 22 million taxpayers who do not file a tax return (Pay As You Earn system)
 - Low attrition rate: results for balanced panel similar to full sample
- Only 11% of taxpayers report positive donations
 - Corner solution issue (aka “censoring”)
 - Potential sample selection bias if giving decision is correlated with sensitivity to price changes

Data descriptives

Figure: Fraction of Donors by Income and Gender



Note: Calculations derived from HMRC's administrative data sources

Figure: Share of Income Donated, by Income and Gender



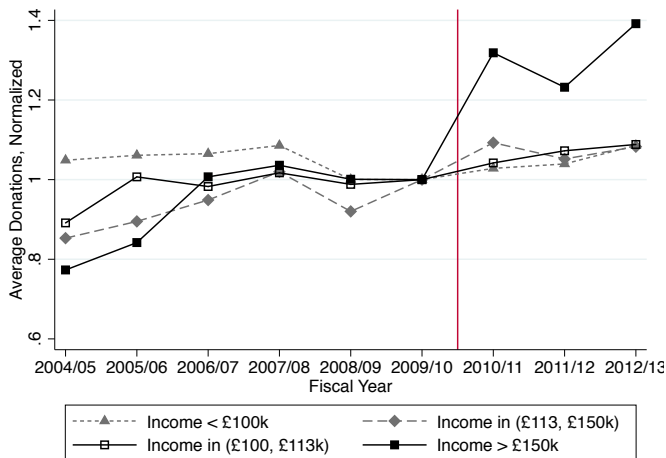
Note: Calculations derived from HMRC's administrative data sources

“Treatment” and “Control” Groups

- Two control groups:
 - Control 1 (C1): $Y_{it} < 100k$
 - Control 2 (C2): $Y_{it} \in (113k, 150k)$
- Two treatment groups:
 - Treatment 1 (T1): $Y_{it} \in (100k, 113k)$
 - Treatment 2 (T2): $Y_{it} > 150k$
- Identification challenges:
 - Pre-reform trends \Rightarrow Graphical analysis
 - Aggregate shocks, e.g. financial crisis \Rightarrow Year fixed effects
 - Time-invariant individual characteristics, e.g. generosity \Rightarrow Individual fixed effects
 - Time-variant individual characteristics, e.g. income \Rightarrow Controls in regression framework

Graphical Diff-in-diff Analysis

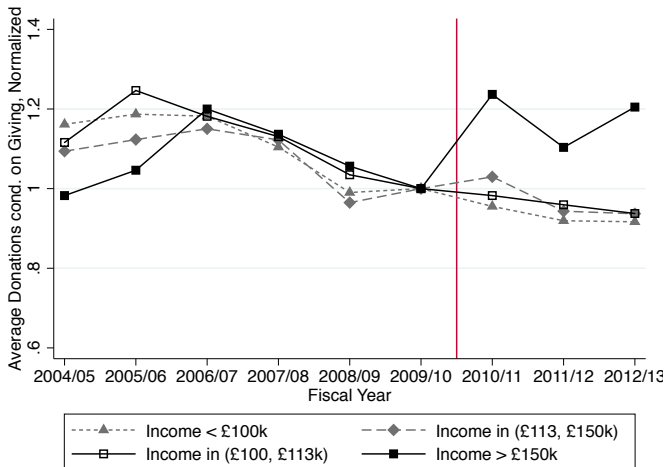
Figure: Normalized Average Donations by Income Group



Note: Calculations derived from HMRC's administrative data sources

Graphical Diff-in-diff Analysis

Figure: Normalized Donations, conditional on Giving



Note: Calculations derived from HMRC's administrative data sources

- 1 Baseline: panel regressions with fixed effects
 - Intensive-margin and extensive-margin elasticities: $\varepsilon_{int}, \varepsilon_{ext}$
 - OLS and IV estimation (first-pound price)
- 2 Differenced regressions with additional IV for Price
 - Intensive-margin elasticity (ε_{int})
 - Avoid price endogeneity problems present in earlier papers, due to income responses to tax changes
- 3 Poisson and negative binomial regressions with fixed effects
 - Estimate total elasticity: $\varepsilon_{total} = \varepsilon_{int} + \varepsilon_{ext}$
 - Allow for zeros in the outcome variable

Baseline Regression Framework

- **Intensive margin:** main regression equation

$$\ln g_{it} = \varepsilon_{INT} \ln p_{it} + \eta_{INT} \ln y_{it} + \delta X_{it} + \alpha_i + \alpha_t + u_{it} \quad (1)$$

- Estimated by OLS or IV only on donors ($g_{it} > 0$)
- α_i, α_t denote individual and year fixed effects
- X_{it} includes age (squared), tax advisor dummy

- **Extensive-margin:** main regression equation

$$D_{it} = \beta \ln p_{it} + \gamma \ln y_{it} + \delta X_{it} + \alpha_i + \alpha_t + v_{it} \quad (2)$$

- Estimated via linear probability model (LPM) on all observations (donors and non-donors, $g_{it} \geq 0$)

Issue #1: Price is Endogenous

- Endogeneity of OLS estimation:
 - Mechanically, $\text{corr}(p_{it}, g_{it}) > 0$, bc large donations can shift taxpayer to a lower tax bracket, increasing the price of giving ($p_{it} = 1 - \tau_{it}$)
 - Upward bias in OLS estimates of ε_{int}
- Standard IV strategy: “First-pound” price of giving
 - Price of giving the taxpayer would face if she made zero donations

$$p_{it}^f = 1 - \tau(y_{it} | g_{it} = 0)$$

- To deal with correlation between giving and income by controlling for
In $y_{it} :=$ disposable income, setting $g_{it} = 0$

Intensive Margin: OLS Regressions (Last-pound price, p)

	Dependent Variable: Log Donations (ln g_{it})			
	(1)	(2)	(3)	(4)
Log Price of Giving	0.055*** (0.004)	0.003 (0.004)	-0.016*** (0.004)	0.016*** (0.004)
Log Disposable Income	0.169*** (0.001)	0.156*** (0.001)	0.154*** (0.001)	0.157*** (0.001)
Individual FE	y	y	y	y
Year FE	n	y	n	y
Other controls	n	n	y	y
Observations	8,275,307	8,275,307	8,240,273	8,240,273
R-squared	0.012	0.043	0.039	0.043
Unique IDs	2,095,064	2,095,064	2,082,159	2,082,159

Note: standard errors in parentheses, clustered at the individual level.

Intensive Margin: IV Regressions ($IV = p^f$)

	Dependent Variable: Log Donations ($\ln g_{it}$)			
	(1)	(2)	(3)	(4)
Log Price of Giving	-0.328*** (0.004)	-0.359*** (0.004)	-0.383*** (0.004)	-0.345*** (0.004)
Log Disposable Income	0.130*** (0.001)	0.118*** (0.001)	0.116*** (0.001)	0.119*** (0.001)
Individual FE	y	y	y	y
Year FE	n	y	n	y
Other controls	n	n	y	y
Observations	7,652,940	7,652,940	7,624,586	7,624,586
R-squared	0.009	0.040	0.036	0.041
Unique IDs	1,472,697	1,472,697	1,466,472	1,466,472

Note: standard errors in parentheses, clustered at the individual level.

Extensive Margin: LPM (Last-pound price, p)

	Dependent Variable: Donor Dummy, $D_{it} \equiv (g_{it} > 0)$			
	(1)	(2)	(3)	(4)
Log Price of Giving	-0.030*** (0.000)	-0.059*** (0.000)	-0.059*** (0.000)	-0.060*** (0.000)
Log Disposable Income	0.010*** (0.000)	0.006*** (0.000)	0.007*** (0.000)	0.007*** (0.000)
<i>Implied Price Elasticity, ϵ_{EXT}</i>	-0.267*** (0.002)	-0.527*** (0.002)	-0.518*** (0.002)	-0.521*** (0.002)
<i>Implied Income Elasticity, η_{EXT}</i>	0.085*** (0.000)	0.057*** (0.000)	0.060*** (0.000)	0.057*** (0.000)
Individual FE	y	y	y	y
Year FE	n	y	n	y
Other controls	n	n	y	y
Observations	73,319,687	73,319,687	71,850,001	71,850,001
Unique IDs	14,149,861	14,149,861	13,700,463	13,700,463
R-squared	0.0548	0.0408	0.0500	0.00979

Note: standard errors in parentheses, clustered at the individual level.

Extensive Margin: LPM ($IV = p^f$)

	Dependent Variable: Donor Dummy, $D_{it} \equiv (g_{it} > 0)$			
	(1)	(2)	(3)	(4)
Log Price of Giving	-0.060*** (0.000)	-0.090*** (0.000)	-0.090*** (0.000)	-0.091*** (0.000)
Log Disposable Income	0.007*** (0.000)	0.004*** (0.000)	0.005*** (0.000)	0.004*** (0.000)
<i>Implied Price Elasticity, ε_{EXT}</i>	-0.533*** (0.002)	-0.801*** (0.002)	-0.789*** (0.002)	-0.794*** (0.002)
<i>Implied Income Elasticity, η_{EXT}</i>	0.065*** (0.000)	0.036*** (0.000)	0.039*** (0.000)	0.036*** (0.000)
Individual FE	y	y	y	y
Year FE	n	y	n	y
Other controls	n	n	y	y
Observations	73,319,687	73,319,687	71,850,001	71,850,001
Unique IDs	14,149,861	14,149,861	13,700,463	13,700,463
R-squared	0.055	0.041	0.050	0.010

Note: standard errors in parentheses, clustered at the individual level.

Issue #2: Earnings Response to Tax Reform

- The tax-price p_{it}^f depends on post-reform earnings (z_{it})
 - Controlling for $\ln y_{it}$ only solves the problem if the relationship is log-linear (unlikely)
- Proposed solution: use lagged value of earnings to construct price IV (similar to Gruber & Saez, 2002). Taking first differences, we use

$$\ln \left(\frac{p_{it}^f(z_{i,t-k})}{p_{i,t-k}^f(z_{i,t-k})} \right) \quad (3)$$

as an instrument for the log change in the first-pound price,

$$\ln \left(\frac{p_{it}^f(z_{i,t})}{p_{i,t-k}^f(z_{i,t-k})} \right) \quad (4)$$

Intensive Margin: Differenced Regressions ($k = 1$)

	Dep. Var.: Log change in Donations, $\ln(g_{it}/g_{i,t-k})$			
	(1)	(2)	(3)	(4)
First Difference ($k = 1$)				
Log change in First-Pound Price	-0.149*** (0.010)	-0.139*** (0.010)	-0.188*** (0.009)	-0.176*** (0.009)
Log change in Disposable Income			0.080*** (0.001)	0.081*** (0.001)
Observations	5,216,321	5,198,174	5,204,515	5,186,411
R-squared R-squared	0.002	0.003	0.006	0.006
Individual FE	y	y	y	y
Year FE	y	y	y	y
Other controls	n	y	n	y

Note: standard errors in parentheses, clustered at the individual level.

Intensive Margin: Differenced Regressions ($k = 2$)

	Dep. Var.: Log change in Donations, $\ln(g_{it}/g_{i,t-k})$			
	(1)	(2)	(3)	(4)
Second Difference ($k = 2$)				
Log change in First-Pound Price	-0.150*** (0.011)	-0.132*** (0.011)	-0.232*** (0.010)	-0.213*** (0.010)
Log change in Disposable Income			0.109*** (0.002)	0.111*** (0.002)
Observations	3,463,375	3,451,745	3,456,133	3,444,530
R-squared	0.003	0.004	0.010	0.012
Individual FE	y	y	y	y
Year FE	y	y	y	y
Other controls	n	y	n	y

Note: standard errors in parentheses, clustered at the individual level.

Intensive Margin: Differenced Regressions ($k = 3$)

	Dep. Var.: Log change in Donations, $\ln(g_{it}/g_{i,t-k})$			
	(1)	(2)	(3)	(4)
Third Difference ($k = 3$)				
Log change in First-Pound Price	-0.205*** (0.015)	-0.176*** (0.016)	-0.378*** (0.013)	-0.355*** (0.013)
Log change in Disposable Income			0.114*** (0.002)	0.116*** (0.002)
Observations	1,955,897	1,949,354	1,951,991	1,945,460
R-squared	0.006	0.007	0.014	0.016
Individual FE	y	y	y	y
Year FE	y	y	y	y
Other controls	n	y	n	y

Note: standard errors in parentheses, clustered at the individual level.

Issue #3: Censoring

- Almost 90% of taxpayers report $g_{it} = 0$
- Could bias our estimated intensive-margin elasticity if decision to donate correlated with donated amount
- Two alternative approaches:
 - Two-step selection model à la Heckman (1979) to estimate ε_{int}
 - Results very similar to differenced model
 - Poisson or Negative binomial regressions to estimate the total elasticity ($\varepsilon_{total} = \varepsilon_{int} + \varepsilon_{ext}$)

Intensive Margin: Two-Step Selection Model

Inverse Mills Ratio (IMR):	Pooled One effect (1)	Pooled Diff effects (2)	Annual One effect (3)	Annual Diff effects (4)
Price Elasticity of Giving	-0.236*** (0.006)	-0.239*** (0.006)	-0.201*** (0.006)	-0.164*** (0.009)
Income Elasticity of Giving	0.139*** (0.002)	0.138*** (0.002)	0.138*** (0.002)	0.136*** (0.002)
P-value on IMR terms	0.000	0.000	0.000	0.000
Observations	5,014,687	5,014,687	5,014,687	5,014,687
R-squared	0.102	0.102	0.100	0.100

Note: robust standard errors in parentheses.

Total Elasticity: Poisson-type Regressions

- Increasingly popular way to model outcomes with a large share of zeros, esp. in trade literature (Santos-Silva & Tenreyro, 2006)
- Allows for linear fixed effects – no incidental parameters problem
- Estimation equation:

$$g_{it} = \exp(\varepsilon \ln p_{it} + \eta \ln y_{it} + \alpha_i + \alpha_t + \delta X_{it}) + u_{it} \quad (5)$$

- where ε can be interpreted as the total price elasticity

Total Elasticity: Poisson Regressions

	Dependent Variable: Donations in Levels (g_{it})			
	(1)	(2)	(3)	(4)
Log First-Pound Price	-1.670*** (0.088)	-1.603*** (0.088)	-1.018*** (0.090)	-0.947*** (0.091)
Log Disposable Income			0.564*** (0.063)	0.561*** (0.036)
Individual FE	y	y	y	y
Year FE	y	y	y	y
Other controls	n	y	n	y
Observations	13,645,910	13,585,847	13,645,910	13,585,847
Unique IDs	1,963,164	1,953,903	1,963,164	1,953,903

Note: standard errors in parentheses, clustered at the individual level.

Total Elasticity: Negative Binomial Regressions

	Dependent Variable: Donations in Levels (g_{it})			
	(1)	(2)	(3)	(4)
Log First-pound Price	-1.573*** (0.004)	-1.798*** (0.005)	-1.080*** (0.006)	-1.161*** (0.006)
Log Disposable Income			0.121*** (0.001)	0.167*** (0.001)
Individual FE	y	y	y	y
Year FE	y	y	y	y
Other controls	n	y	n	y
Observations	13,645,910	13,585,847	13,645,910	13,585,847
Unique IDs	1,963,164	1,953,903	1,963,164	1,953,903

Note: standard errors in parentheses, clustered at the individual level.

Heterogeneous Elasticities

- One interesting question is whether the elasticities vary by income level (or age, gender)
- We construct income groups based on *average* income during the period 2005-2012
- Results:
 - ε_{int} larger (in abs. value) for high-income taxpayers
 - ε_{ext} larger for low-income taxpayers
 - $\varepsilon_{total} \approx -0.8$ for high-income and $\varepsilon_{total} \approx -1.6$ for low-income

Heterogeneous Elasticities by Income: Intensive Margin

Income group	Dependent Variable: Log change in Donations, $\ln(g_{it}/g_{i,t-k})$				
	$p0 - p25$ (1)	$p25 - p50$ (2)	$p50 - p75$ (3)	$p75 - p95$ (4)	$p95 - p100$ (5)
Log change in First-Pound Price	0.089 (0.065)	-0.048 (0.043)	-0.055** (0.025)	-0.098*** (0.013)	-0.220*** (0.028)
Log change in Disposable Income	0.045*** (0.004)	0.077*** (0.004)	0.088*** (0.003)	0.100*** (0.003)	0.114*** (0.004)
Individual FE	y	y	y	y	y
Year FE	y	y	y	y	y
Other controls	y	y	y	y	y
Observations	100,089	526,510	1,483,141	2,167,162	909,509
R-squared	0.005	0.007	0.006	0.007	0.007

Note: standard errors in parentheses, clustered at the individual level.

Heterogeneous Elasticities by Income: Extensive Margin

	Dependent Variable: Donor Dummy $I(g_{it} > 0)$				
	$p0 - p25$	$p25 - p50$	$p50 - p75$	$p75 - p95$	$p95 - p100$
	(1)	(2)	(3)	(4)	(5)
<i>Implied Price Elasticity, ε_{EXT}</i>	-1.583*** (0.018)	-0.998*** (0.010)	-0.455*** (0.005)	-0.270*** (0.002)	-0.170*** (0.004)
<i>Implied Income Elasticity, η_{EXT}</i>	0.091*** (0.002)	0.092*** (0.001)	0.079*** (0.001)	0.075*** (0.001)	0.076*** (0.001)
Individual FE	y	y	y	y	y
Year FE	y	y	y	y	y
Other controls	y	y	y	y	y
Observations	13,772,160	18,005,842	19,684,814	15,780,001	4,607,184
Unique IDs	3,385,342	3,422,862	3,434,745	2,757,835	699,679
R-squared	0.002	0.006	0.010	0.022	0.037

Note: standard errors in parentheses, clustered at the individual level.

Heterogeneous Elasticities by Income: Total Elasticity

	Dependent Variable: Log Donations ($\ln g_{it}$)				
	$p0 - p25$	$p25 - p50$	$p50 - p75$	$p75 - p95$	$p95 - p100$
Negative Binomial	(1)	(2)	(3)	(4)	(5)
Log Price of Giving	-1.852*** (0.027)	-1.624*** (0.022)	-0.719*** (0.013)	-0.736*** (0.009)	-0.745*** (0.011)
Log Disposable Income	0.085*** (0.003)	0.243*** (0.004)	0.295*** (0.003)	0.252*** (0.002)	0.076*** (0.003)
Poisson					
Log Price of Giving	-2.216*** (0.099)	-1.120*** (0.053)	-0.514*** (0.042)	-0.413*** (0.024)	-1.208*** (0.173)
Log Disposable Income	0.164*** (0.016)	0.413*** (0.018)	0.437*** (0.023)	0.531*** (0.016)	0.582*** (0.044)
Individual FE	y	y	y	y	y
Year FE	y	y	y	y	y
Other controls	y	y	y	y	y
Observations	754,910	1,970,915	3,847,126	4,932,990	2,079,906
Unique IDs	125,025	296,276	561,094	695,029	276,479

Note: standard errors in parentheses, clustered at the individual level.

Heterog. Elasticities by Gender & Age: Intensive Margin

	Dep. Var.: Change in Log Donations, $\ln g_{it} / \ln g_{i,t-k}$				
	Men (1)	Women (2)	Age < 40 (3)	Age 40 – 65 (4)	Age > 65 (5)
Log change in First-Pound Price	-0.192*** (0.011)	-0.140*** (0.017)	-0.204*** (0.033)	-0.226*** (0.013)	-0.118*** (0.014)
Log change in Disposable Income	0.082*** (0.002)	0.081*** (0.002)	0.082*** (0.004)	0.072*** (0.002)	0.104*** (0.004)
Individual FE	y	y	y	y	y
Year FE	y	y	y	y	y
Other controls	y	y	y	y	y
Observations	3,358,795	1,827,616	588,690	2,711,995	1,885,726
R-squared	0.007	0.006	0.010	0.006	0.006

Note: standard errors in parentheses, clustered at the individual level.

Heterog. Elasticities by Gender & Age: Extensive Margin

	Dependent Variable: Donor Dummy $I(g_{it} > 0)$				
	Men (1)	Women (2)	Age < 40 (3)	Age 40 – 65 (4)	Age > 65 (5)
<i>Implied Price Elasticity, ε_{EXT}</i>	-0.653*** (0.002)	-0.724*** (0.004)	-1.273*** (0.007)	-0.625*** (0.003)	-0.263*** (0.004)
<i>Implied Income Elasticity, η_{EXT}</i>	0.041*** (0.000)	0.046*** (0.001)	0.078*** (0.001)	0.046*** (0.000)	0.071*** (0.001)
Individual FE	y	y	y	y	y
Year FE	y	y	y	y	y
Other controls	y	y	y	y	y
Observations	47,406,495	24,443,506	20,581,542	39,550,708	11,717,751
Unique IDs	8,905,195	4,795,268	5,789,633	8,003,184	2,467,174
R-squared	0.011	0.014	0.014	0.011	0.007

Note: standard errors in parentheses, clustered at the individual level.

Heterog. Elasticities by Gender & Age: Total Elasticity

	Dependent Variable: Log Donations ($\ln g_{it}$)				
	Men	Women	Age < 40	Age 40 – 65	Age > 65
Negative Binomial	(1)	(2)	(3)	(4)	(5)
Log Price of Giving	-1.070*** (0.006)	-0.879*** (0.010)	-1.101*** (0.012)	-0.959*** (0.009)	-0.891*** (0.006)
Log Disposable Income	0.153*** (0.001)	0.250*** (0.003)	0.157*** (0.002)	0.179*** (0.002)	0.123*** (0.002)
Poisson					
Log Price of Giving	-1.036*** (0.115)	-0.356*** (0.118)	-1.324*** (0.320)	-0.867*** (0.102)	-0.900*** (0.195)
Log Disposable Income	0.540*** (0.039)	0.634*** (0.072)	0.601*** (0.108)	0.536*** (0.040)	0.531*** (0.140)
Individual FE	y	y	y	y	y
Year FE	y	y	y	y	y
Other controls	y	y	y	y	y
Observations	8,497,798	5,088,049	2,307,553	7,068,175	3,462,729
Unique IDs	1,201,588	752,315	446,348	1,113,730	560,224

Note: standard errors in parentheses, clustered at the individual level.

Summary of Elasticity Estimates

- Panel fixed-effects (OLS/IV):
 - Intensive-margin price elasticity: $\varepsilon^{int} \in (-0.23, -0.17)$
 - Extensive-margin price elasticity: $\varepsilon^{ext} \in (-0.81, -0.78)$
 - Total price elasticity ($\varepsilon_{ext} + \varepsilon_{int}$) ≈ -1
 - With some heterogeneity across income levels
- Total price elasticity (Poisson/Negative Binomial):
 - Total price elasticity: $\varepsilon \in (-1.16, -0.94)$
 - Consistent with int+ext margin estimates

Interpretation of Results: Optimal Subsidy?

- Traditionally, focus on the “unit-elasticity” rule:
 - If $|\varepsilon| > 1$, increase the subsidy
 - If $|\varepsilon| < 1$, decrease the subsidy
- This criterion assumes away:
 - Individual's utilities
 - Crowding-out of private donations if there is public contribution
 - (Positive) external effects of donations

Unit-Elasticity Rule: “Treasury Efficiency”

Define the net benefit (B) of a subsidy to charitable donations as:

$$B = (1 - s) DN$$

where $D = \textit{per capita}$ donation, $N = \textit{number of donors}$, $s = \textit{subsidy}$

The subsidy is “treasury efficient” (assuming no crowding in/out) if:

$$\frac{\partial B}{\partial s} = -DN + \left(\frac{\partial D}{\partial s} N + \frac{\partial N}{\partial s} D \right) (1 - s) > 0$$

$$\left(\frac{\partial D}{\partial (1 - s)} \frac{(1 - s)}{D} + \frac{\partial N}{\partial (1 - s)} \frac{(1 - s)}{N} \right) < -1$$

$$(\varepsilon_{int} + \varepsilon_{ext}) < -1$$

Optimal Subsidy: Theoretical Framework

- We extend Saez's (2004) optimal tax expenditures framework:
 - Allow for extensive-margin responses
 - Let the government place a different value on private donations vs direct govt subsidies
- Our elasticity estimates are only consistent with the current subsidy being optimal if the govt values private donations less than direct govt provision
 - “Merit goods” argument: govt believes that donors' preferences do not line up with the majority's preferences
 - This result holds even when accounting for private donors' “warm-glow” utility in the social welfare function